

Lecture 8.

3.3. Stein's method of exchangeable pairs. The exchangeable pairs approach within Stein's method was first used in the paper [Dia77] which, however, attributed the method to Charles Stein. Later, this technique was presented in a systematic way in Stein's monograph [St86]. We recall that a pair (X, X') of random elements on a common probability space is said to be *exchangeable*, if (X, X') has the same distribution as (X', X) . In the book [St86, Lecture III], it is highlighted that a given real random variable W is close in distribution to a standard normal variable N , whenever one can construct an exchangeable pair (W, W') such that W' is close to W in *some* sense and such that the *linear regression property*

$$\mathbb{E}[W' - W \mid \mathcal{G}] = -\lambda W \quad (3.6)$$

is satisfied for some small $\lambda > 0$ and $\text{Var}\left(\frac{1}{2\lambda}\mathbb{E}[(W' - W)^2 \mid \mathcal{G}]\right)$ is small, where $\mathcal{G}$ is a σ -algebra that contains $\sigma\{W\}$. Note the relation (3.6) forces W to be centered.

Theorem 3.4 (C. Stein, 1986). *Let (W, W') be an exchangeable pair of random variables defined on a common probability space $(\Omega, \mathcal{F}, \mathbb{P})$ such that $\text{Var}(W) = 1$, (3.6) holds for some $\lambda > 0$ and $W \in L^3(\mathbb{P})$, then we have*

$$d_W(W, N) \leq \sqrt{2/\pi} \sqrt{\text{Var}\left(\frac{1}{2\lambda}\mathbb{E}[(W - W')^2 \mid \mathcal{G}]\right)} + \frac{1}{2\lambda}\mathbb{E}[|W - W'|^3],$$

where $N \sim \mathcal{N}(0, 1)$.

Proof. Let $f : \mathbb{R} \rightarrow \mathbb{R}$ belong to $C^2(\mathbb{R})$ with $\|f\|_\infty \leq 2$, $\|f'\|_\infty \leq \sqrt{2/\pi}$ and $\|f''\|_\infty \leq 2$, then it is clear that $f(W')W$ and $f(W)W$ are integrable. It follows from the exchangeability that $\mathbb{E}[(W - W')f(W)] = \mathbb{E}[(W' - W)f(W')]$, from which we get $\mathbb{E}[(W - W')[f(W) + f(W')]] = 0$. So

$$\begin{aligned} 0 &= \mathbb{E}[(W - W')[-f(W) + f(W')]] + 2\mathbb{E}[(W - W')f(W)] \\ &= \mathbb{E}[(W - W')[-f(W) + f(W')]] + 2\mathbb{E}[f(W)\mathbb{E}(W - W' \mid \mathcal{G})] \\ &= \mathbb{E}[(W - W')[-f(W) + f(W')]] + 2\lambda\mathbb{E}[f(W)W] \quad \text{by (3.6)} \end{aligned}$$

that is, $\mathbb{E}[(W' - W)[f(W') - f(W)]] = 2\lambda\mathbb{E}[f(W)W]$; and particularly, $\mathbb{E}[(W' - W)^2] = 2\lambda$. Therefore,

$$\begin{aligned} 2\lambda\mathbb{E}[f(W)W - f'(W)] &= \mathbb{E}\left((W' - W)[f(W') - f(W)] - 2\lambda f'(W)\right) \\ &= \mathbb{E}\left(f'(W)[(W' - W)^2 - 2\lambda]\right) + \mathbb{E}\left[(W' - W)^3 \int_0^1 (1-t)f''(W + t(W' - W))dt\right] \\ &= \mathbb{E}\left[f'(W)\left(\mathbb{E}[(W' - W)^2 \mid \mathcal{G}] - 2\lambda\right)\right] + \mathbb{E}\left[(W' - W)^3 \int_0^1 (1-t)f''(W + t(W' - W))dt\right], \end{aligned}$$

which together with Proposition 3.3 gives us the desired bound. $\square$

In the following, we will apply Stein's bound of exchangeable pairs to obtain a rate of convergence for the CLT under the assumption of finite fourth moment.

Proposition 3.5 (Quantitative CLT via exchangeable pairs). *Let $\{X_i\}_i$ be i.i.d. with mean zero and variance one such that $\mathbb{E}[|X_1|^4] < \infty$. Then with $V_n = \frac{1}{\sqrt{n}}(X_1 + \dots + X_n)$,*

$$d_W(V_n, \mathcal{N}(0, 1)) \leq \frac{C}{\sqrt{n}},$$

where $C = \left(\frac{\mathbb{E}[X_1^4]-1}{2\pi}\right)^{\frac{1}{2}} + 4\mathbb{E}[|X_1|^3]$.

Proof. Let us first construct the exchangeable pair:

$$V'_n = V_n + \frac{X'_I - X_I}{\sqrt{n}},$$

where $\{X'_i\}_i$ is an independent copy of $\{X_i\}_i$, and I is a uniform random variable with values in $\{1, \dots, n\}$ such that I is independent from $\{X_i\}_i$ and $\{X'_i\}_i$. We claim that V_n and V'_n are exchangeable. Indeed, for any $t, s \in \mathbb{R}$, we have

$$\begin{aligned} \mathbb{P}(V_n \leq t, V'_n \leq s) &= \sum_{k=1}^n \mathbb{P}(V_n \leq t, V'_n \leq s, I = k) \\ &= \sum_{k=1}^n \mathbb{P}\left(U_k + \frac{X_k}{\sqrt{n}} \leq t, U_k + \frac{X'_k}{\sqrt{n}} \leq s, I = k\right) \quad \text{with } U_k = V_n - \frac{X_k}{\sqrt{n}} \\ &= \sum_{k=1}^n \frac{1}{n} \mathbb{P}\left(U_k + \frac{X_k}{\sqrt{n}} \leq t, U_k + \frac{X'_k}{\sqrt{n}} \leq s\right) \\ &= \sum_{k=1}^n \frac{1}{n} \mathbb{P}\left(U_k + \frac{X'_k}{\sqrt{n}} \leq t, U_k + \frac{X_k}{\sqrt{n}} \leq s\right) \end{aligned}$$

by the independence among U_k, X_k, X'_k . Then, we have

$$\begin{aligned} \mathbb{P}(V_n \leq t, V'_n \leq s) &= \sum_{k=1}^n \mathbb{P}\left(U_k + \frac{X'_k}{\sqrt{n}} \leq t, U_k + \frac{X_k}{\sqrt{n}} \leq s, I = k\right) \\ &= \mathbb{P}(V'_n \leq t, V_n \leq s). \end{aligned}$$

That is, $(V_n, V'_n) \stackrel{(\text{law})}{=} (V'_n, V_n)$. Next, we write, with $\mathcal{G} = \sigma\{X_1, \dots, X_n\}$,

$$\begin{aligned} \mathbb{E}[V'_n - V_n | \mathcal{G}] &= \frac{1}{\sqrt{n}} \mathbb{E}[X'_I - X_I | \mathcal{G}] = \frac{1}{\sqrt{n}} \sum_{k=1}^n \mathbb{E}[(X'_k - X_k) \mathbf{1}_{\{I=k\}} | \mathcal{G}] \\ &= \frac{1}{\sqrt{n}} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[X'_k - X_k | \mathcal{G}] \quad \text{as } I \text{ is independent from } \mathcal{G} \\ &= -\frac{1}{n} \frac{1}{\sqrt{n}} \sum_{k=1}^n X_k \quad \text{since } \mathbb{E}[X'_k | \mathcal{G}] = 0 \\ &= -\frac{1}{n} V_n. \end{aligned}$$

That is, (3.6) holds true with $(W, W', \lambda) = (V_n, V'_n, \frac{1}{n})$.

Next, by similar arguments, we can have

$$\begin{aligned} \mathbb{E}[(V_n - V'_n)^2 | \mathcal{G}] &= \frac{1}{n} \mathbb{E}[(X_I - X'_I)^2 | \mathcal{G}] = \frac{1}{n^2} \sum_{k=1}^n \mathbb{E}[(X_k - X'_k)^2 | \mathcal{G}] \\ &= \frac{1}{n^2} \sum_{k=1}^n \mathbb{E}[|X_k|^2 - 2X_k X'_k + |X'_k|^2 | \mathcal{G}] \\ &= \frac{1}{n^2} \sum_{k=1}^n (X_k^2 + 1). \end{aligned}$$

Thus,

$$\text{Var} \left(\frac{1}{2\lambda} \mathbb{E}[(V_n - V'_n)^2 | \mathcal{G}] \right) = \text{Var} \left(\frac{1}{2n} \sum_{k=1}^n (1 + X_k^2) \right) = \frac{1}{4n} (\mathbb{E}[X_1^4] - 1).$$

Finally, let us look at the third moment of the difference $V'_n - V_n$:

$$\begin{aligned} \frac{1}{2\lambda} \mathbb{E}[|V'_n - V_n|^3] &= \frac{1}{2\sqrt{n}} \mathbb{E}[|X_I - X'_I|^3] = \frac{1}{2\sqrt{n}} \sum_{k=1}^n \frac{1}{n} \mathbb{E}[|X_k - X'_k|^3] \\ &= \frac{1}{2\sqrt{n}} \mathbb{E}[|X_1 - X'_1|^3] \leq \frac{1}{2\sqrt{n}} \mathbb{E}[4|X_1|^3 + 4|X'_1|^3] \\ &= \frac{4}{\sqrt{n}} \mathbb{E}[|X_1|^3]. \end{aligned}$$

Combining the above bounds, we can get

$$d_W(V_n, \mathcal{N}(0, 1)) \leq \left\{ \left(\frac{\mathbb{E}[X_1^4] - 1}{2\pi} \right)^{\frac{1}{2}} + 4\mathbb{E}[|X_1|^3] \right\} n^{-\frac{1}{2}}.$$

Hence, the proof is completed. $\square$

3.4. 2nd-order Gaussian Poincaré inequality. The Gaussian Poincaré inequality (Proposition 2.5) indicates that the variance of $F = f(\mathbf{G})$, with $f \in \mathbb{D}^{1,2}(\mathbb{R}^d)$ and $\mathbf{G} \sim \mathcal{N}_d(0, \mathbf{I})$, is controlled by the average magnitude of DF (i.e., $\mathbb{E}[\|DF\|_{\mathbb{R}^d}^2]$). That is, when the latter is small, the random variable F has small variance, i.e., small fluctuation from its mean. The bound in Proposition 2.5 becomes an exact equality only when $F = f(\mathbf{G})$ is a linear combination of the components $G_1, \dots, G_d$, that is, when F is a Gaussian random variable. Naively, if f is close to an affine function, then F would be expected to be close to a Gaussian distribution (with matched mean and variance). Then, if we have a good control over the “second-order” derivatives of f (i.e., $\text{Hess}f(\mathbf{G})$), we should have a proximity of F to a Gaussian. It is S. Chatterjee who first came up with this idea and made it rigorous via Stein’s method; see [Cha09]. Let us illustrate this idea more with the quadratic form. Let

$$F = \langle \mathbf{G}, A\mathbf{G} \rangle_{\mathbb{R}^d} = \sum_{i,j=1}^d A_{ij} G_i G_j,$$

where $A = (A_{ij})$ is a symmetric $d \times d$ matrix. By the orthogonal decomposition of a real symmetric matrix, we can write

$$A = Q^T \Lambda Q,$$

where $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_d)$ is the diagonal matrix with entries being the eigenvalues of A , Q is an orthogonal matrix (i.e., $QQ^T = \mathbf{I}$). Then, we can rewrite $F = \sum_{i=1}^d \lambda_i Y_i^2$, with $(Y_1, \dots, Y_d) = Q\mathbf{G}$ again a standard Gaussian vector. By the CLT heuristic, F is approximately Gaussian if no eigenvalue dominates in the sum, or $\max \lambda_i^2$ is much smaller than $\sum_i \lambda_i^2$. Note that for $f(x) = \langle x, Ax \rangle_{\mathbb{R}^d}$, we have the Hessian matrix

$$\text{Hess}f(x) = 2A \quad \text{so that the operator norm } \|\text{Hess}f(x)\|_{\text{op}} = 2 \max_i |\lambda_i|,$$

while $\text{Var}(F) = 2 \sum_{i=1}^d \lambda_i^2$. This leads to the heuristic that F is approximately Gaussian if the typical value of $\|\text{Hess}f(x)\|_{\text{op}}^2$ is much smaller than the variance of F .

Theorem 3.6 (2nd-order GPI, [Cha09]). Suppose $f \in C^2(\mathbb{R}^d)$ and $F = f(\mathbf{G})$ with $\mathbf{G} \sim \mathcal{N}_d(0, \mathbf{I})$. Suppose F has zero mean and variance $\sigma^2 \in (0, \infty)$. Then, we have

$$d_{\text{TV}}(F, \mathcal{N}(0, \sigma^2)) \leq \frac{3}{\sigma^2} \mathbb{E} \left[\|\text{Hess} f(\mathbf{G})\|_{\text{op}}^4 \right]^{\frac{1}{4}} \mathbb{E} \left[\|\nabla f(\mathbf{G})\|_{\mathbb{R}^d}^4 \right]^{\frac{1}{4}}. \quad (3.7)$$

Proof. By the definition of total-variation distance, we have $d_{\text{TV}}(F, \mathcal{N}(0, \sigma^2)) = d_{\text{TV}}(\frac{F}{\sigma}, \mathcal{N}(0, 1))$. Then, it is enough to consider the case where $\text{Var}(F) = \sigma^2 = 1$. According to Stein's method, it suffices to prove a uniform bound for the expectation

$$\mathbb{E}[F\varphi(F)] - \mathbb{E}[\varphi'(F)] \quad \text{with } \|\varphi'\|_{\infty} \leq 2.$$

In the rest of the proof, we assume that the RHS of (3.7) is finite, otherwise the proof is trivial. With F centered, we can write $F = \mathcal{L}\mathcal{L}^{-1}F$, and therefore with $\mathcal{L} = -\delta\nabla$ ($\nabla = D$),

$$\begin{aligned} \mathbb{E}[F\varphi(F)] &= \mathbb{E}[-\delta\nabla\mathcal{L}^{-1}F\varphi(F)] = \mathbb{E}[\langle -\nabla\mathcal{L}^{-1}F, \nabla\varphi(F) \rangle_{\mathbb{R}^d}] \quad \text{by duality relation} \\ &= \mathbb{E}[\langle -\nabla\mathcal{L}^{-1}F, \nabla F \rangle_{\mathbb{R}^d} \varphi'(F)] \quad \text{by chain rule,} \end{aligned}$$

and in particular, we have $1 = \mathbb{E}[F^2] = \mathbb{E}[\langle -\nabla\mathcal{L}^{-1}F, \nabla F \rangle_{\mathbb{R}^d}]$. Thus, the expression $\mathbb{E}[F\varphi(F)] - \mathbb{E}[\varphi'(F)]$ is uniformly bounded by $2\sqrt{\text{Var}(\langle -\nabla\mathcal{L}^{-1}F, \nabla F \rangle_{\mathbb{R}^d})}$. From previous sections, we know that

$$-\nabla\mathcal{L}^{-1}F = \int_0^\infty e^{-t} P_t \nabla F dt \quad \text{and} \quad A \cdot (-\nabla\mathcal{L}^{-1}F) = \int_0^\infty e^{-t} A \cdot (P_t \nabla F) dt; \quad (3.8)$$

by the same arguments, we can get (with $\nabla^2 f = \text{Hess} f \in \mathbb{R}^{d \times d}$)

$$-\nabla^2\mathcal{L}^{-1}F = \int_0^\infty e^{-2t} P_t \nabla^2 F dt \quad \text{and} \quad (-\nabla^2\mathcal{L}^{-1}F) \cdot y = \int_0^\infty e^{-2t} (P_t \nabla^2 F) \cdot y dt, \quad (3.9)$$

where $A \cdot y$ means that the matrix $A \in \mathbb{R}^{d \times d}$ acts on the vector $y \in \mathbb{R}^d$,

Applying the Gaussian Poincaré inequality, we have

$$\begin{aligned} \sqrt{\text{Var}(\langle -\nabla\mathcal{L}^{-1}F, \nabla F \rangle_{\mathbb{R}^d})} &\leq \sqrt{\mathbb{E}[\|\nabla \langle -\nabla\mathcal{L}^{-1}F, \nabla F \rangle_{\mathbb{R}^d}\|_{\mathbb{R}^d}^2]} \\ &\leq \sqrt{\mathbb{E}[\|(-\nabla^2\mathcal{L}^{-1}F) \cdot (\nabla F)\|_{\mathbb{R}^d}^2]} + \sqrt{\mathbb{E}[\|(\nabla^2 F) \cdot (-\nabla\mathcal{L}^{-1}F)\|_{\mathbb{R}^d}^2]}, \end{aligned} \quad (3.10)$$

where in the last step, we applied Minkowski's inequality. Now let us deal with the first term in (3.10): using (3.9) with Minkowski's inequality, we write

$$\begin{aligned} \left\| \|(-\nabla^2\mathcal{L}^{-1}F) \cdot (\nabla F)\|_{\mathbb{R}^d} \right\|_{L^2(\mathbb{P})} &\leq \left\| \int_0^\infty e^{-2t} \|P_t \nabla^2 F\|_{\text{op}} \|\nabla F\|_{\mathbb{R}^d} dt \right\|_{L^2(\mathbb{P})} \\ &\leq \int_0^\infty e^{-2t} \left\| \|P_t \nabla^2 F\|_{\text{op}} \|\nabla F\|_{\mathbb{R}^d} \right\|_{L^2(\mathbb{P})} dt \\ &\leq \int_0^\infty e^{-2t} \left\| \|P_t \nabla^2 F\|_{\text{op}} \right\|_{L^4(\mathbb{P})} \left\| \|\nabla F\|_{\mathbb{R}^d} \right\|_{L^4(\mathbb{P})} dt \\ &\leq \int_0^\infty e^{-2t} \left\| \|\nabla^2 F\|_{\text{op}} \right\|_{L^4(\mathbb{P})} \left\| \|\nabla F\|_{\mathbb{R}^d} \right\|_{L^4(\mathbb{P})} dt \quad (*) \\ &= \frac{1}{2} \left\| \|\nabla^2 F\|_{\text{op}} \right\|_{L^4(\mathbb{P})} \left\| \|\nabla F\|_{\mathbb{R}^d} \right\|_{L^4(\mathbb{P})}, \end{aligned}$$

where we used the contraction property of the O-U semigroup in the step (*). The second term in (3.10) can be dealt with in the same way, so that the proof is concluded. $\square$

Exercise 3.7. Work out the details for the step (*) and the details for the second term in (3.10).